



**QUALIFICATION: Bachelor of Science in Applied Mathematics and Statistics**

<b>QUALIFICATION CODE:</b> 07BSAM	<b>LEVEL:</b> 7
<b>COURSE CODE:</b> AEM702S	<b>COURSE NAME:</b> Applied Econometrics Modelling
<b>SESSION:</b> NOVEMBER 2024	<b>PAPER:</b> THEORY
<b>DURATION:</b> 3 HOURS	<b>MARKS:</b> 100

**FIRST OPPORTUNITY EXAMINATION QUESTION PAPER**

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**INSTRUCTIONS**

1. There are 7 questions, answer ALL the questions by showing all the necessary steps.
2. Write clearly and neatly.
3. Number the answers clearly.
4. Round your answers to at least four decimal places, if applicable.

**PERMISSIBLE MATERIALS**

1. Nonprogrammable scientific calculators with no cover.

**THIS QUESTION PAPER CONSISTS OF 3 PAGES (Including this front page)**

**ATTACHMENTS**

Four statistical distribution tables (t-, z-,  $\chi^2$ - and F-distribution tables)

### Question 1 [11 Marks]

- 1.1. Compare and contrast econometrics, economic theory, mathematical economics and economic statistics [4]
- 1.2. Briefly discuss the problem of multicollinearity in multiple linear regression model. Your discussion should include definition, two consequences, two methods of detections and two possible remedial measures. [7]

### Question 2 [20 Marks]

2. Consider a two-variable linear regression model  $Y_i = \beta_1 + \beta_2 X_i + u_i$  for  $i = 1, 2, 3, \dots, n$
- 2.1. Show that the OLS estimator  $\hat{\beta}_2$  is an unbiased estimator of  $\beta_2$ . Hint  $\hat{\beta}_2 = \frac{\sum x_i Y_i}{\sum x_i^2}$ , where  $x_i = X_i - \bar{X}$  [6]
- 2.2. Show that the  $Var(\hat{\beta}_2) = \frac{\sigma^2}{\sum x_i^2}$  [6]
- 2.3. Show that  $\hat{\beta}_2$  is the best estimator of  $\beta_2$ . Hint: Consider an alternative linear unbiased estimator  $\hat{\beta}_2^* = \sum w_i Y_i$  of  $\beta_2$  and show that  $Var(\hat{\beta}_2^*) \geq Var(\hat{\beta}_2)$ . [8]

### Question 3 [10 Marks]

- 3.1. Consider the general ( $k$ -variable) linear regression model

$$\begin{matrix} \mathbf{y} & = & \mathbf{X} & \boldsymbol{\beta} & + & \mathbf{u} \\ n \times 1 & & n \times k & k \times 1 & & n \times 1 \end{matrix}$$

Show that an OLS estimator  $\hat{\boldsymbol{\beta}} = (\mathbf{X}' \mathbf{X})^{-1} \mathbf{X}' \mathbf{y}$  is an unbiased estimator of  $\boldsymbol{\beta}$ . [5]

- 3.2. Suppose a researcher collected data on personal saving ( $S$ ) and personal income ( $I$ ) for 31-year period and fitted a linear regression model

$$S_i = \beta_1 + \beta_2 I_i + u_i$$

Assume that a graphical inspection suggests that  $u_i$ 's are heteroscedastic so that a researcher employed the **Goldfeld Quandt test** by removing  $c = 9$  central observations after arranging the data based on income. Applying OLS to each subset, a researcher obtained the following results.

For subset I:  $\hat{S}_{1i} = -738.84 + 0.008I_i$  With  $RSS_1 = \sum \hat{u}_{1i}^2 = 144,771.5$

For Subset II:  $\hat{S}_{2i} = 1141.07 + 0.029I_i$  With  $RSS_2 = \sum \hat{u}_{2i}^2 = 769,899.2$

Based on the above results, test if there is any evidence of heteroscedasticity at 5% level of significance. [5]

### Question 4 [21 Marks]

4. A marketing department is interested in the effects of changing advertising levels for television and internet on sales ( $Y$ ). They vary  $X_2$ =total expenditure on TV advertisement in \$, and  $X_3$ =total expenditure on internet advertisement in \$. Answer the following questions based on the summary of sample values:

$$n = 20; \quad \bar{Y} = 183.8186; \quad \mathbf{y}' \mathbf{y} = 686084.6; \quad RSS = 608.6247; \quad TSS = 10299.09$$

$$\mathbf{X}' \mathbf{X} = \begin{pmatrix} 20 & 416.5343 & 406.487 \\ 416.5343 & 9546.826 & 8733.245 \\ 46.487 & 8733.245 & 9111.308 \end{pmatrix}$$

$$\mathbf{X}' \mathbf{y} = \begin{pmatrix} 3676.373 \\ 78940.14 \\ 77022.41 \end{pmatrix} \quad (\mathbf{X}' \mathbf{X})^{-1} = \begin{pmatrix} 0.800494 & -0.01832 & -0.01815 \\ -0.01832 & 0.00127 & -0.0004 \\ -0.01815 & -0.0004 & 0.001303 \end{pmatrix}$$

$$Var - cov(\hat{\boldsymbol{\beta}}) = \begin{pmatrix} 28.65885 & -0.65588 & -0.6498 \\ -0.65588 & 0.045468 & -0.01432 \\ -0.6498 & -0.01432 & 0.046649 \end{pmatrix}$$

- 4.1. Compute the point estimator of  $\beta = \begin{pmatrix} \beta_1 \\ \beta_2 \\ \beta_3 \end{pmatrix}$  and interpret each partial coefficient [4]
- 4.2. Construct ANOVA table and test the hypothesis  $H_0: \beta_2 = \beta_3 = 0$  at 5% level of sign. [5]
- 4.3. Compute an unbiased estimate of the residual variance [2]
- 4.4. Test the significance of the partial regression coefficient for advertisement expense on TV ( $X_2$ ),  $\beta_2$ . Use 5% level of significancy. [4]
- 4.5. Compute and interpret the coefficient of multiple determination. [3]
- 4.6. If a marketing department decided to invest \$10 on TV advertisement and \$3 on internet advertisement, then what will be the predicted sells for such investment based on the fitted model? [3]

#### Question 5 [6 Marks]

5. Consider the following infinite distributed lag model:

$$Y_t = \alpha + \beta_0 X_t + \beta_1 X_{t-1} + \beta_2 X_{t-2} + \beta_3 X_{t-3} + \dots + u_t \quad \text{for } 0 < \lambda \leq 1$$

Show how the Koyck transformation can be used to produce the following type of model

$$Y_t = \alpha(1 - \lambda) + \beta_0 X_t + \lambda Y_{t-1} + v_t \quad [6]$$

#### Question 6 [20 Marks]

6. Consider the following regression result for expenditure on new plant and equipment ( $Y$ ) on sales ( $X$ ) in billions of dollars and lagged value of  $Y$ .

$$\begin{aligned} \hat{Y}_t &= -15.104 + 0.629X_t + 0.272Y_{t-1} \\ se &= (4.7294) \quad (0.0978) \quad (0.1148) \\ d &= 1.5185, \quad \text{durbin } h = 1.3403 \end{aligned}$$

Answer the following questions based on this result.

- 6.1. If we assume that this model resulted from a Koyck-type transformation, then

- 6.1.1. What is the estimate for rate of decline or decay this model? [2]  
 6.1.2. Compute the median lag [2]  
 6.1.3. Compute the mean lag [2]  
 6.1.4. What is the short-run or impact multiplier value for this model? Provide interpretation of the value as well. [2]  
 6.1.5. Compute the estimate for the coefficient of the first lag,  $X_{t-1}$ . Hint: Use the Koyck scheme. [2]

- 6.2. Assuming that

$$Y_t^* = \alpha + \beta_0 X_t + u_t,$$

where  $Y^*$  = desired, or long-run, expenditure for new plant and equipment.

Derive the partial adjustment model.

- 6.2.1. Compute the coefficient of partial adjustment [2]  
 6.2.2. Estimate the parameters of this model [4]

#### Question 7 [12 Marks]

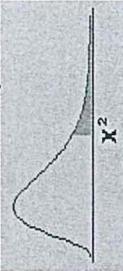
7. Consider the following structural equations

$$\begin{aligned} Y_{1t} &= \beta_{10} + \beta_{12}Y_{2t} + \gamma_{11}X_{1t} + u_{1t} \\ Y_{2t} &= \beta_{20} + \beta_{21}Y_{1t} + u_{2t} \end{aligned}$$

- 7.1. Derive the reduced form equations expressed in the form of  $Y_{1t}$  and  $Y_{2t}$ . [10]  
 7.2. Determine which of the preceding equations are identified (either just or over). [2]

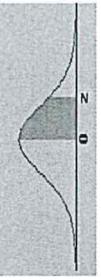
==== END OF QUESTION PAPER====

Right tail areas for the Chi-square Distribution



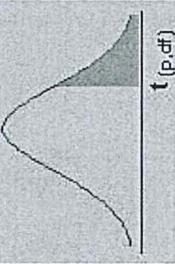
dfarea	.995	.990	.975	.950	.900	.750	.500	.250	.100	.050	.025	.010	.005
1	0.00004	0.00016	0.00098	0.00393	0.01579	0.10153	0.45494	1.32330	2.70554	3.84146	5.02389	6.63490	7.87944
2	0.01003	0.02010	0.05064	0.10259	0.21072	0.57536	1.38629	2.77259	4.60517	5.99146	7.37776	9.21034	10.59663
3	0.07172	0.11483	0.21580	0.35185	0.58437	1.21253	2.36597	4.10834	6.25139	7.81473	9.34840	11.34487	12.83816
4	0.20699	0.29711	0.48442	0.71072	1.06362	1.92256	3.35669	5.38527	7.77944	9.48773	11.14229	13.27670	14.86026
5	0.41174	0.55430	0.83121	1.14548	1.61031	2.67460	4.35146	6.62568	9.23636	11.07050	12.83250	15.08627	16.74960
6	0.67573	0.87209	1.23734	1.63538	2.20413	3.45460	5.34812	7.84080	10.64464	12.59159	14.44938	16.81189	18.54758
7	0.98926	1.23904	1.68987	2.16735	2.83311	4.23485	6.34581	9.03715	12.01704	14.06714	16.01276	18.47531	20.27774
8	1.34441	1.64650	2.17973	2.73264	3.48954	5.07064	7.34412	10.21885	13.36157	15.50731	17.53455	20.09024	21.95495
9	1.73493	2.08790	2.70039	3.32511	4.16816	5.89883	8.34283	11.38875	14.68366	16.91898	19.02277	21.66599	23.58935
10	2.15586	2.55821	3.24697	3.94030	4.86518	6.73720	9.34182	12.54886	15.98718	18.30704	20.48318	23.20925	25.18818
11	2.60322	3.05348	3.81575	4.57481	5.57778	7.58414	10.34100	13.70069	17.27501	19.67514	21.92005	24.72497	26.75685
12	3.07382	3.57057	4.40379	5.22603	6.30380	8.43842	11.34032	14.84540	18.54935	21.02607	23.33666	26.21697	28.29952
13	3.56503	4.10692	5.00875	5.89186	7.04150	9.29907	12.33976	15.98391	19.81193	22.36203	24.73560	27.68825	29.81947
14	4.07467	4.66043	5.62873	6.57063	7.78953	10.16531	13.33927	17.11693	21.06414	23.68479	26.11895	29.14124	31.31935
15	4.60092	5.22935	6.26214	7.26094	8.54676	11.03654	14.33886	18.24509	22.30713	24.99579	27.48839	30.57791	32.80132
16	5.14221	5.81221	6.90766	7.96165	9.31224	11.91222	15.33850	19.36886	23.54183	26.29623	28.84535	31.99993	34.26719
17	5.69722	6.40776	7.56419	8.67176	10.08519	12.79193	16.33818	20.48868	24.76904	27.58711	30.19101	33.40866	35.71847
18	6.26480	7.01491	8.23075	9.39046	10.86494	13.67529	17.33790	21.60489	25.98942	28.86930	31.52638	34.80531	37.15645
19	6.84397	7.63273	8.90632	10.11701	11.65091	14.56200	18.33765	22.71781	27.20357	30.14353	32.85233	36.19087	38.58226
20	7.43384	8.26040	9.59078	10.85081	12.44261	15.45177	19.33743	23.82769	28.41198	31.41043	34.16961	37.56623	39.99685
21	8.03365	8.89720	10.28290	11.59131	13.23960	16.34438	20.33723	24.93478	29.61509	32.67057	35.47888	38.93217	41.40106
22	8.64272	9.54249	10.98232	12.33801	14.04149	17.23962	21.33704	26.03927	30.81328	33.92444	36.78071	40.28936	42.79565
23	9.26042	10.19572	11.68855	13.09051	14.84796	18.13730	22.33688	27.14134	32.00690	35.17246	38.07563	41.63840	44.18128
24	9.88623	10.85636	12.40115	13.84843	15.65868	19.03725	23.33673	28.24115	33.19624	36.41503	39.36408	42.97982	45.55851
25	10.51965	11.52398	13.11972	14.61141	16.47341	19.93934	24.33659	29.33885	34.38159	37.63248	40.64647	44.31410	46.92789
26	11.16024	12.19815	13.84390	15.37916	17.29188	20.84343	25.33646	30.43457	35.56317	38.88514	41.92317	45.64168	48.28988
27	11.80759	12.87850	14.57338	16.15140	18.11390	21.74940	26.33634	31.52841	36.74122	40.11327	43.19451	46.96294	49.64492
28	12.46134	13.56471	15.30786	16.92788	18.93924	22.65716	27.33623	32.62049	37.91592	41.33714	44.46079	48.27824	50.99338
29	13.12115	14.25645	16.04707	17.70837	19.76774	23.56659	28.33613	33.71091	39.08747	42.55697	45.72229	49.58788	52.33562
30	13.78672	14.95346	16.79077	18.49266	20.59923	24.47761	29.33603	34.79974	40.25602	43.77297	46.97924	50.89218	53.67196

Area between 0 and z



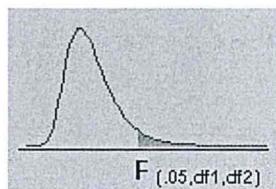
	<b>0.00</b>	<b>0.01</b>	<b>0.02</b>	<b>0.03</b>	<b>0.04</b>	<b>0.05</b>	<b>0.06</b>	<b>0.07</b>	<b>0.08</b>	<b>0.09</b>
<b>0.0</b>	0.0000	0.0040	0.0080	0.0120	0.0160	0.0199	0.0239	0.0279	0.0319	0.0359
<b>0.1</b>	0.0398	0.0438	0.0478	0.0517	0.0557	0.0596	0.0636	0.0675	0.0714	0.0753
<b>0.2</b>	0.0793	0.0832	0.0871	0.0910	0.0948	0.0987	0.1026	0.1064	0.1103	0.1141
<b>0.3</b>	0.1179	0.1217	0.1255	0.1293	0.1331	0.1368	0.1406	0.1443	0.1480	0.1517
<b>0.4</b>	0.1554	0.1591	0.1628	0.1664	0.1700	0.1736	0.1772	0.1808	0.1844	0.1879
<b>0.5</b>	0.1915	0.1950	0.1985	0.2019	0.2054	0.2088	0.2123	0.2157	0.2190	0.2224
<b>0.6</b>	0.2257	0.2291	0.2324	0.2357	0.2389	0.2422	0.2454	0.2486	0.2517	0.2549
<b>0.7</b>	0.2580	0.2611	0.2642	0.2673	0.2704	0.2734	0.2764	0.2794	0.2823	0.2852
<b>0.8</b>	0.2881	0.2910	0.2939	0.2967	0.2995	0.3023	0.3051	0.3078	0.3106	0.3133
<b>0.9</b>	0.3159	0.3186	0.3212	0.3238	0.3264	0.3289	0.3315	0.3340	0.3365	0.3389
<b>1.0</b>	0.3413	0.3438	0.3461	0.3485	0.3508	0.3531	0.3554	0.3577	0.3599	0.3621
<b>1.1</b>	0.3643	0.3665	0.3686	0.3708	0.3729	0.3749	0.3770	0.3790	0.3810	0.3830
<b>1.2</b>	0.3849	0.3869	0.3888	0.3907	0.3925	0.3944	0.3962	0.3980	0.3997	0.4015
<b>1.3</b>	0.4032	0.4049	0.4066	0.4082	0.4099	0.4115	0.4131	0.4147	0.4162	0.4177
<b>1.4</b>	0.4192	0.4207	0.4222	0.4236	0.4251	0.4265	0.4279	0.4292	0.4306	0.4319
<b>1.5</b>	0.4332	0.4345	0.4357	0.4370	0.4382	0.4394	0.4406	0.4418	0.4429	0.4441
<b>1.6</b>	0.4452	0.4463	0.4474	0.4484	0.4495	0.4505	0.4515	0.4525	0.4535	0.4545
<b>1.7</b>	0.4554	0.4564	0.4573	0.4582	0.4591	0.4599	0.4608	0.4616	0.4625	0.4633
<b>1.8</b>	0.4641	0.4649	0.4656	0.4664	0.4671	0.4678	0.4686	0.4693	0.4699	0.4706
<b>1.9</b>	0.4713	0.4719	0.4726	0.4732	0.4738	0.4744	0.4750	0.4756	0.4761	0.4767
<b>2.0</b>	0.4772	0.4778	0.4783	0.4788	0.4793	0.4798	0.4803	0.4808	0.4812	0.4817
<b>2.1</b>	0.4821	0.4826	0.4830	0.4834	0.4838	0.4842	0.4846	0.4850	0.4854	0.4857
<b>2.2</b>	0.4861	0.4864	0.4868	0.4871	0.4875	0.4878	0.4881	0.4884	0.4887	0.4890
<b>2.3</b>	0.4893	0.4896	0.4898	0.4901	0.4904	0.4906	0.4909	0.4911	0.4913	0.4916
<b>2.4</b>	0.4918	0.4920	0.4922	0.4925	0.4927	0.4929	0.4931	0.4932	0.4934	0.4936
<b>2.5</b>	0.4938	0.4940	0.4941	0.4943	0.4945	0.4946	0.4948	0.4949	0.4951	0.4952
<b>2.6</b>	0.4953	0.4955	0.4956	0.4957	0.4959	0.4960	0.4961	0.4962	0.4963	0.4964
<b>2.7</b>	0.4965	0.4966	0.4967	0.4968	0.4969	0.4970	0.4971	0.4972	0.4973	0.4974
<b>2.8</b>	0.4974	0.4975	0.4976	0.4977	0.4977	0.4978	0.4979	0.4980	0.4981	
<b>2.9</b>	0.4981	0.4982	0.4982	0.4983	0.4984	0.4984	0.4985	0.4986	0.4986	
<b>3.0</b>	0.4987	0.4987	0.4987	0.4988	0.4988	0.4988	0.4989	0.4989	0.4990	0.4990

t table with right tail probabilities



df	p	0.40	0.25	0.10	0.05	0.025	0.01	0.005	0.0005	t <sub>(p, df)</sub>
1	0.324920	1.000000	3.077684	6.313752	12.70620	31.82052	63.65674	636.6192		
2	0.288675	0.816497	1.885618	2.919986	4.30265	6.96456	9.92484	31.5991		
3	0.276671	0.764892	1.637744	2.353363	3.18245	4.54070	5.84091	12.9240		
4	0.270722	0.740697	1.533206	2.131847	2.77645	3.74695	4.60409	8.6103		
5	0.267181	0.7266687	1.475884	2.015048	2.57058	3.36493	4.03214	6.8688		
6	0.264835	0.717558	1.439756	1.943180	2.44691	3.14267	3.70743	5.9588		
7	0.263167	0.711142	1.414924	1.894579	2.36462	2.99795	3.49948	5.4079		
8	0.261921	0.7065387	1.396815	1.859548	2.30600	2.89646	3.35539	5.0413		
9	0.260955	0.702722	1.383029	1.833113	2.26216	2.82144	3.24984	4.7809		
10	0.260185	0.699812	1.372184	1.812461	2.22814	2.76377	3.16927	4.5869		
11	0.259556	0.697445	1.363430	1.795885	2.20099	2.71808	3.10581	4.4370		
12	0.259033	0.695483	1.356217	1.782288	2.17881	2.68100	3.05454	4.3178		
13	0.258591	0.693829	1.350171	1.770933	2.16037	2.65031	3.01228	4.2208		
14	0.258213	0.692417	1.345030	1.761310	2.14479	2.62449	2.97684	4.1405		
15	0.257885	0.691197	1.340606	1.753050	2.13145	2.60248	2.94671	4.0728		
16	0.257599	0.690132	1.336757	1.745884	2.11991	2.58349	2.92078	4.0150		
17	0.257347	0.689195	1.333379	1.739607	2.10982	2.56693	2.89823	3.9651		
18	0.257123	0.6888364	1.330391	1.734064	2.10092	2.55238	2.87844	3.9216		
19	0.256923	0.687621	1.327728	1.729133	2.09302	2.53948	2.86093	3.8834		
20	0.256743	0.686954	1.325341	1.724718	2.08596	2.52798	2.84534	3.8495		
21	0.256580	0.686552	1.323188	1.720743	2.07961	2.51765	2.83136	3.8193		
22	0.256432	0.685805	1.321237	1.717144	2.07587	2.50832	2.81876	3.7921		
23	0.256297	0.684850	1.317836	1.710882	2.06866	2.49987	2.80734	3.7676		
24	0.256173	0.684430	1.319460	1.713872	2.06390	2.49216	2.79694	3.7454		
25	0.256060	0.684430	1.316345	1.708141	2.05954	2.48511	2.78744	3.7251		
26	0.255955	0.684043	1.314972	1.705618	2.05553	2.47863	2.77871	3.7066		
27	0.255858	0.683685	1.313703	1.703288	2.05183	2.47266	2.77068	3.6896		
28	0.255768	0.683353	1.312527	1.701131	2.04841	2.46714	2.76326	3.6739		
29	0.255684	0.683044	1.311434	1.699127	2.04523	2.46202	2.75639	3.6594		
30	0.255605	0.682756	1.310415	1.697261	2.04227	2.45726	2.75000	3.6460		

Table for  $\alpha=.05$



$df2 \backslash df1$	1	2	3	4	5	6	7	8	9	10	12
1	161.448	199.500	215.707	224.583	230.162	233.986	236.768	238.883	240.543	241.882	243.906
2	18.513	19.000	19.164	19.247	19.296	19.329	19.353	19.371	19.384	19.396	19.413
3	10.128	9.552	9.277	9.117	9.014	8.941	8.887	8.845	8.812	8.786	8.745
4	7.709	6.944	6.591	6.388	6.256	6.163	6.0942	6.041	5.998	5.964	5.912
5	6.608	5.786	5.409	5.192	5.050	4.950	4.876	4.818	4.772	4.735	4.678
6	5.987	5.143	4.757	4.533	4.387	4.284	4.207	4.147	4.099	4.060	3.999
7	5.591	4.737	4.347	4.120	3.972	3.866	3.787	3.726	3.676	3.637	3.575
8	5.318	4.459	4.066	3.838	3.688	3.581	3.501	3.438	3.388	3.347	3.284
$df2$	9	5.117	4.256	3.863	3.633	3.482	3.374	3.293	3.229	3.178	3.137
	10	4.965	4.103	3.708	3.478	3.326	3.217	3.136	3.072	3.020	2.978
	11	4.844	3.982	3.587	3.358	3.204	3.095	3.012	2.948	2.896	2.854
	12	4.747	3.885	3.490	3.259	3.106	2.996	2.913	2.849	2.796	2.753
	13	4.667	3.806	3.411	3.179	3.025	2.915	2.832	2.767	2.714	2.671
	14	4.600	3.739	3.344	3.112	2.958	2.848	2.764	2.699	2.645	2.602
	15	4.543	3.682	3.287	3.056	2.901	2.791	2.707	2.641	2.587	2.544
	16	4.494	3.634	3.239	3.007	2.852	2.741	2.657	2.591	2.537	2.494
	17	4.451	3.591	3.197	2.965	2.810	2.699	2.614	2.548	2.494	2.450
	18	4.414	3.555	3.160	2.928	2.773	2.661	2.577	2.510	2.456	2.412
	19	4.381	3.522	3.127	2.895	2.740	2.628	2.544	2.477	2.423	2.378
	20	4.351	3.493	3.098	2.866	2.711	2.599	2.514	2.441	2.393	2.348